

Tentamen Kansrekening – Open book

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1. Denote by X the number of fixed points of a random permutation of n objects.

Compute $E(X)$, the variance, and $E(X^3)$!

Hint: Use computations with indicator functions!

Solution: Write $X = \sum_{i=1}^n 1_{A_i}$ where A_i is the event that i is a fixed point.

We have for the probabilities

$$\begin{aligned}P(A_1) &= \frac{1}{n} \\P(A_1 \cap A_2) &= \frac{1}{n(n-1)} \\P(A_1 \cap A_2 \cap A_3) &= \frac{1}{n(n-1)(n-2)}\end{aligned}$$

So we get for the moments

$$E(X) = nE(1_{A_1}) = 1$$

and

$$E(X^2) = (n^2 - n)P(A_1 \cap A_2) + nP(A_1) = 2$$

So the variance is 1. Finally,

$$\begin{aligned}E(X^3) &= n(n-1)(n-2)P(A_1 \cap A_2 \cap A_3) + 3n(n-1)P(A_1 \cap A_2) + nP(A_1) \\&= 1 + 3 + 1 = 5\end{aligned}$$

Note the $3n(n-1)$ in the second term on the r.h.s. We can check the prefactors when we realize that $n(n-1)(n-2) + 3n(n-1) + n = n^3$. This is clear because they account for all possible choices of three indices between 1 and n .

2. Suppose that A and B are two events with $P(A) = \frac{3}{4}$ and $P(B) = \frac{1}{3}$. Show that always the inequality $\frac{1}{12} \leq P(A \cap B) \leq \frac{1}{3}$ holds

Solution: The r.h.s. is clear by considering B . The l.h.s. can be seen by

$$P(A \cap B) = P(A) + P(B) - P(A \cup B) \geq P(A) + P(B) - 1$$

substituting the explicit numbers.

3. Be N a random variable that is distributed according to Poisson with parameter λ .

a) Compute the exponential moment generating function $E(e^{tN})$.

Solution: well known from lecture and book, $E(e^{tN}) = \exp \lambda(e^t - 1)$

b) Now take a sequence of i.i.d. random Normal variables with expected value μ and variance σ^2 and show that $E(e^{tX_i}) = e^{\frac{t^2\sigma^2}{2} + t\mu}$

Next we define the random variable $S := \sum_{i=1}^N X_i$, where the random variables N and X_i are as above.

c) Compute the exponential moment generating function $E(e^{tS})$!

Solution:

$$\begin{aligned} E(e^{tN}) &= \sum_{n=0}^{\infty} P(N = n) E(e^{t \sum_{i=1}^n X_i}) \\ &= e^{-\lambda} \sum_{n=0}^{\infty} \frac{\lambda^n}{n!} E(e^{tX_1})^n \\ &= \exp \lambda(e^{\frac{t^2\sigma^2}{2} + t\mu} - 1) \end{aligned}$$

d) Compute $E(S)$ and the variance of S !

Hint for d): Use c)! However, if you don't like c), you can also avoid it in the computation

Solution: $E(S) = \lambda\mu$

$$E(S^2) = \lambda^2\mu^2 + \lambda(\sigma^2 + \mu^2)$$

$$E(S^2) - E(S)^2 = \lambda(\sigma^2 + \mu^2)$$

One can avoid part c) in the computation, if you use the formula for the expected value and variance of a sum with a random number of terms that was treated in the exercises.

4. Be X an exponential random variable with parameter 1.

(a) Compute the density of $\log X$

Solution: For the distribution function we have $P(\log X \leq x) = 1 - e^{-e^x}$.

Taking derivatives gives the density $e^{-e^x} e^x$.

Be Y another exponential random variable with parameter 1, independent of X

(b) Compute the joint density of the random vector $(X, \frac{Y}{X+Y})$.

Hint: Use the multidimensional transformation formula, and note that $\frac{Y}{X+Y}$ takes values only in the interval $[0, 1]$

Solution: Put $u = x$ and $v = \frac{y}{x+y}$. Then the inverse transformation is given by $x = u$ and $y = \frac{vu}{1-v}$. The functional determinant is given by $\frac{u}{(1-v)^2}$.

So the joint density is

$$\exp\left(-\frac{u}{1-v}\right) \frac{u}{(1-v)^2}$$

for $0 < u < \infty$ and $0 < v < 1$.

(c) Compute the density of $\frac{Y}{X+Y}$.

Solution: as we see from (b) by integration over u it is uniform on the interval $[0, 1]$. More precisely, the marginal distribution on the v -variable is obtained by integration over u . Making a substitution $\tilde{u} = \frac{u}{1-v}$ this integral becomes

$$\int_0^\infty \exp\left(-\frac{u}{1-v}\right) \frac{u}{(1-v)^2} du = \int_0^\infty \exp(-\tilde{u}) \tilde{u} d\tilde{u} = 1$$

That the last integral is 1 can be seen without looking at computations or tables from the fact that it has become v -independent because the remaining quantity must be a probability density on $[0, 1]$.

5. Be X_n a sequence of independent geometric random variables with n -dependent parameter $p_n = 1 - e^{-n}$. (We use notations according to the table in KaDe page 251, meetkundig)

Use the Borel-Cantelli Lemma to show that X_n converges to 0 almost surely! Is the assumption of independence really necessary to reach this conclusion?

Solution: That X_n converges to zero almost surely means equivalently that only finitely many non-zeros occur. Now, this is implied by the convergence of the series

$$\sum_{n=1}^{\infty} P(X_n \neq 0) = \sum_{n=1}^{\infty} (1 - p_n) = \sum_{n=1}^{\infty} e^{-n} < \infty$$

For this conclusion we don't need the independence.